

# A Process Algebraic Framework for Estimating the Energy Consumption in Ad-hoc Wireless Sensor Networks\*

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## ABSTRACT

We present a framework for modelling ad-hoc Wireless Sensor Networks (WSNs) and studying both their connectivity properties and their performances in terms of energy consumption, throughput and other relevant indices. Our framework is based on a probabilistic process calculus where system executions are driven by Markovian probabilistic schedulers, allowing us to translate process terms into discrete time Markov chains (DTMCs) and use the probabilistic model checker PRISM to automatically evaluate/estimate the connectivity properties and the energy costs of the networks. To the best of our knowledge, this is the first work that proposes a unique framework for studying qualitative (e.g., by proving the equivalence of components or the correctness of a behaviour) and quantitative aspects of WSNs using a tool that allows both exact and approximate (via Monte Carlo simulation) analyses. We demonstrate our framework at work by considering different communication strategies based on gossip routing protocols, for a typical topology and a mobility scenario.

## Categories and Subject Descriptors

C.4 [Computer Systems Organization]: Performance of systems—*modeling techniques*; C.2.1 [Computer-Communication Networks]: Network architecture and design—*Wireless Communication*

## General Terms

Theory, reliability

## Keywords

Process algebra, simulation, sensor networks

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## 1. INTRODUCTION

Wireless Sensor Networks (WSNs) [2] are collections of spatially distributed sensing devices equipped with limited computing and radio communication capabilities. They are employed in a variety of applications, ranging from military surveillance to health care or assisted living, and from smart cities to environmental monitoring. A typical sensor network is characterized by a large number of sensor nodes, which are densely deployed, and have frequent topology changes due to the mobility of its devices. Nodes communicate using wireless transmission in a specified range, with communication between nodes implemented in terms of routing protocols. A critical issue in wireless sensor networks is the limited availability of energy within the network devices. Therefore, judicious choice of routing protocols that can reduce the nodes' power consumption is crucial, not only for the performance of each single node, but also for the throughput of the network lifetime. In this paper we introduce a framework for the specification, modelling and automated analysis or simulation of connectivity properties and the evaluation or estimation of energy consumption in ad-hoc WSNs. The framework is based on a variant of the Probabilistic Energy-aware Broadcast Unicast and Multicast (PEBUM) calculus introduced in [5], and supports the automatic performance evaluation through (approximate) probabilistic model checking, e.g., using the PRISM model checker [9]. The advantages of using a process algebra rely on its compositional nature which allows us to decompose both the model construction and the qualitative analysis. Concerning the quantitative analysis we take advantage of the various kind of properties which can be specified in terms of a temporal logic and verified using the model checking technique. The proposed calculus is built around nodes, representing the sensor devices of the systems, and locations, identifying the position cells across which each device may move inside the network. Node mobility is governed by probability distributions. Instead, wireless synchronisations are controlled by sequential processes inside the nodes: each transmission broadcasts a message within a given transmission range. The semantics of our calculus is inspired by Segala's probabilistic automata [15] driven by schedulers to resolve the nondeterministic choice among the probability distributions over target states. Differently from [5], in this work we assume that nodes are not equipped with a unique identifier and they all share the same transmission frequency. These choices reflect the fact that transmissions in ad-hoc sensor networks are directed to a geographical location rather than to a specific node and, due to the low-cost hardware of sensors, only one frequency is used

Networks		Processes	
$M, N ::= \mathbf{0}$	Empty network	$P, Q ::= \mathbf{0}$	Inactive process
$  [P]_l^{\mathbf{J}}$	Sensor Node	$  (\tilde{x}).P$	Input
$  M_1   M_2$	Parallel composition	$  \langle \tilde{w} \rangle_{L,r}.P$	Output
		$  [w_1 = w_2]P, Q$	Matching
		$  A\langle \tilde{w} \rangle$	Recursion

Table 1: Syntax

at a given time [17]. Moreover, in contrast to [6, 5], in this paper we employ *Makovian probabilistic schedulers*, mapping the non-deterministic choices among the different actions a system may enable into probability distributions. As a consequence, the labelled transition system underlying process terms is a discrete time Markov chain (DTMC), which can be used for automatically performing a range of qualitative and quantitative analyses by means of probabilistic model checking, e.g., using PRISM. We define a probabilistic observational congruence in the style of [13] to equate networks exhibiting the same connectivity behaviour. As in [5], and in contrast to [12], the notion of observability is associated with specific locations in the network reflecting the fact that in ad-hoc WSNs the transmissions are not addressed to specific nodes but to specific locations. We provide a coinductive characterization of the observational congruence based on a probabilistic labelled bisimilarity. Finally, we define an energy-aware preorder over networks, to contrast networks having the same behaviour, but different energy costs. We use our framework for a comparative study of gossip based routing protocols for wireless ad-hoc sensor networks. We address the problem of the state space explosion using the statistical model checking implemented in PRISM. The paper is organised as follows. Section 2 presents the calculus, its observational semantics expressed in terms of behavioural equivalences and a characterization of it based on a notion of probabilistic bisimilarity. In Section 3 an energy-aware preorder over networks is defined: it allows us to compare the average energy cost of different networks but exhibiting the same connectivity behaviour. Finally, in Section 4, gossip based routing protocols for different scenarios are considered and we show the framework at work studying the sensitivity of the performances of these protocols to some configuration parameters. Finally, Section 5 discusses the related bibliography and concludes the paper.

## 2. A CALCULUS FOR WSN

We present a variant of the PEBUM calculus presented in [5] which focuses on the main features of ad-hoc wireless sensor networks. Specifically, nodes are not equipped with a unique identifier and only one transmission frequency is used.

*Syntax.* We use letters  $l$  for *locations*,  $r$  for *transmission radii*,  $x$  and  $y$  for *variables*. *Closed values* contain locations, transmission radii and any basic value (booleans, integers, ...). *Values* include also variables. We use  $u$  and  $v$  for closed values and  $w$  for (open) values. We write  $\tilde{v}, \tilde{w}$  for tuples of values. We write  $\mathcal{N}$  for the set of all networks and  $\mathbf{Loc}$  for the set of all locations. While movements may be assumed to be continuous, we identify locations as the countable set of cells that constitute the observing areas within the network. The syntax of our calculus is shown in Table 1. This

is defined in a two-level structure: the lower one for processes, the upper one for networks. Networks are collections of sensor nodes running in parallel and communicating messages. As usual,  $\mathbf{0}$  denotes the empty network and  $M_1 | M_2$  denotes the parallel composition of two networks. We denote by  $\prod_{i \in I} M_i$  the parallel composition of the networks  $M_i$ , for  $i \in I$ .  $[P]_l^{\mathbf{J}}$  denotes a sensor node located at the physical location  $l$  and executing the process  $P$ .  $\mathbf{J}$  is the transition matrix of a discrete time Markov chain modelling node mobility: each entry  $\mathbf{J}_{lk}$  is the probability that the sensor node located at  $l$  moves to the location  $k$ . Hence,  $\sum_{k \in \mathbf{Loc}} \mathbf{J}_{lk} = 1$  for all locations  $l \in \mathbf{Loc}$ . Static nodes inside a network are associated with the identity Markov chain, i.e., the identity matrix  $\mathbf{J}_{ll} = 1$  for all  $l \in \mathbf{Loc}$  and  $\mathbf{J}_{lk} = 0$  for all  $l \neq k$ . Processes are sequential and live within the nodes:  $\mathbf{0}$  is the inactive process;  $(\tilde{x}).P$  is ready to listen to a transmission, while  $\langle \tilde{w} \rangle_{L,r}.P$  is ready to transmit. In  $(\tilde{x}).P$ , the variables in  $\tilde{x}$  are bound with scope in  $P$ . As to the output form, the tag  $r$  represents the transmission radius of the sender, while the tag  $L$  is used to maintain the set of physical locations of the intended recipients:  $L = \mathbf{Loc}$  represents a broadcast transmission, while a finite set of locations  $L$  denotes a multicast communication (unicast if  $L$  is a singleton). As stated in the introduction, communication protocols for ad-hoc sensor networks are usually intended to reach a certain location, rather than a specific device, due to the absence of global identifiers associated with the sensor devices. The remaining syntactic forms are standard:  $[w_1 = w_2]P, Q$  behaves as  $P$  if  $w_1 = w_2$ , and as  $Q$  otherwise.  $A\langle \tilde{w} \rangle$  is the process defined via a (possibly recursive) definition  $A\langle \tilde{x} \rangle \stackrel{\text{def}}{=} P$ , with  $|\tilde{x}| = |\tilde{w}|$  where  $\tilde{x}$  contains all variables appearing free in  $P$ .

*Probability distributions for networks.* We denote by  $\mu_l^{\mathbf{J}}$  the probability distribution associated with a node located at  $l$  with transition matrix  $\mathbf{J}$ , i.e., the function over  $\mathbf{Loc}$  such that  $\mu_l^{\mathbf{J}}(k) = \mathbf{J}_{lk}$  for all  $k \in \mathbf{Loc}$ . Let  $M$  be a network. We denote by  $M\{[P]_k^{\mathbf{J}}/[P]_l^{\mathbf{J}}\}$  the network obtained by replacing  $l$  with  $k$  inside the sensor node  $[P]_l^{\mathbf{J}}$ . We also denote by  $\llbracket M \rrbracket_{\mu_l^{\mathbf{J}}}$  the probability distribution over networks induced by  $\mu_l^{\mathbf{J}}$  and defined by: for all networks  $M'$ ,

$$\llbracket M \rrbracket_{\mu_l^{\mathbf{J}}}(M') = \begin{cases} \mu_l^{\mathbf{J}}(k) & \text{if } M' = M\{[P]_k^{\mathbf{J}}/[P]_l^{\mathbf{J}}\} \\ 0 & \text{otherwise.} \end{cases}$$

Intuitively,  $\llbracket M \rrbracket_{\mu_l^{\mathbf{J}}}(M')$  is the probability that the network  $M$  evolves to  $M'$  due to the movement of the sensor node  $[P]_l^{\mathbf{J}}$ . We say that  $M'$  is in the support of  $\llbracket M \rrbracket_{\mu_l^{\mathbf{J}}}$  if  $\llbracket M \rrbracket_{\mu_l^{\mathbf{J}}}(M') \neq 0$ . We write  $\llbracket M \rrbracket_{\Delta}$  for the Dirac distribution on the network  $M$ , i.e., the probability distribution defined as:  $\llbracket M \rrbracket_{\Delta}(M) = 1$  and  $\llbracket M \rrbracket_{\Delta}(M') = 0$  for all  $M' \neq M$ . Finally, we let  $\theta$  range over  $\{\mu_l^{\mathbf{J}} \mid \mathbf{J} \text{ is a transition matrix and } l \in \mathbf{Loc}\} \cup \{\Delta\}$ .

(R-Bcast) $\frac{[(\tilde{v})_{L,r}.P]_l^J \mid \prod_{i \in I} [(\tilde{x}_i).P_i]_{l_i}^{J_i} \rightarrow \llbracket [P]_l^J \mid \prod_{i \in I} [P_i\{\tilde{v}/\tilde{x}_i\}]_{l_i}^{J_i} \rrbracket_\Delta}{\llbracket [P]_l^J \mid \prod_{i \in I} [P_i\{\tilde{v}/\tilde{x}_i\}]_{l_i}^{J_i} \rrbracket_\Delta}$	where $\forall i \in I. d(l, l_i) \leq r$ and $ \tilde{x}_i  =  \tilde{v} $	
(R-Move) $\frac{[P]_l^J \rightarrow \llbracket [P]_l^J \rrbracket_{\mu_l^J}}{\llbracket [P]_l^J \rrbracket_{\mu_l^J}}$	(R-Par) $\frac{M \rightarrow \llbracket M' \rrbracket_\theta}{M N \rightarrow \llbracket M' N \rrbracket_\theta}$	(R-Struct) $\frac{N \equiv M \quad M \rightarrow \llbracket M' \rrbracket_\theta \quad M' \equiv N'}{N \rightarrow \llbracket N' \rrbracket_\theta}$

**Table 2: Reduction Semantics**

*Reduction semantics.* The dynamics of the calculus is specified by the *probabilistic reduction relation* ( $\rightarrow$ ) described in Table 2: it takes the form  $M \rightarrow \llbracket M' \rrbracket_\theta$  denoting a transition that leaves from  $M$  and leads to a probability distribution  $\llbracket M' \rrbracket_\theta$ . As usual, reduction relies on structural congruence ( $\equiv$ ), such that, e.g.,  $M|N \equiv N|M$ ,  $(M|N)|M' \equiv M|(N|M')$  and  $M|\mathbf{0} \equiv M$ . Nodes cannot be created or destroyed, and move autonomously. Node connectivity is verified by looking at the physical location and the transmission radius of the sender: a message broadcast by a node is received only by the nodes that lie in the area delimited by the transmission radius of the sender. We presuppose a function  $d(\cdot, \cdot)$  which returns the distance between two locations.

Rule (R-Bcast) models the transmission of a tuple of messages  $\tilde{v}$  by a sensor node located at  $l$  and using a radius  $r$ . The index set  $I$  may be empty, i.e., the rule can be applied even if no nodes are ready to receive. The radius  $r$  associated with the output action denotes the transmission radius of that communication which may depend on the energy consumption strategy adopted by the surrounding protocol. All the nodes that lie in the range of the sender (i.e., such that  $d(l, l_i) \leq r$ ) will receive the messages. Rule (R-Move) deals with node mobility: a node  $[P]_l^J$  executing a move action will reach a location with a probability described by the distribution  $\mu_l^J$  that depends on the Markov chain  $\mathbf{J}$  statically associated with the node. The remaining rules are standard. Given a network  $M$ , we write  $M \rightarrow_\theta N$  if  $M \rightarrow \llbracket M' \rrbracket_\theta$  and  $N$  is in the support of  $\llbracket M' \rrbracket_\theta$ . Following [6], an execution for  $M$  is a (possibly infinite) sequence of steps  $M \rightarrow_{\theta_1} M_1 \rightarrow_{\theta_2} M_2 \dots$

*Observational Semantics.* According to a standard practice, we formalise the observational semantics of our calculus in terms of a notion of *barb* that provides the basic unit of observation [13]. As in other calculi for wireless communication, the definition of barb is naturally expressed in terms of message transmission. We denote by  $behave(M) = \{\llbracket M' \rrbracket_\theta \mid M \rightarrow \llbracket M' \rrbracket_\theta\}$  the set of the possible behaviours of  $M$ . In order to solve the nondeterminism in a network execution, we consider each possible probabilistic transition  $M \rightarrow \llbracket M' \rrbracket_\theta$  as arising from a *probabilistic scheduler* defined as follows.

**DEFINITION 1 (SCHEDULER).** A probabilistic scheduler is a total function  $F$  assigning to a network  $M$  a distribution  $\phi$  on the set  $behave(M)$ .

We denote by *Sched* the set of all probabilistic schedulers. Given a network  $M$  and a scheduler  $F$ , we define the set of all executions starting from  $M$  and driven by  $F$  as:

$$Exec_M^F = \{e = M_0 \rightarrow_{p_1 \theta_1} M_1 \rightarrow_{p_2 \theta_2} M_2 \dots \mid M_0 \equiv M \text{ and} \\ \forall j > 0 : M_{j-1} \rightarrow \llbracket M'_j \rrbracket_{\theta_j}, p_j = F(M_{j-1})(\llbracket M'_j \rrbracket_{\theta_j}) \\ \text{and } M_j \text{ is in the support of } \llbracket M'_j \rrbracket_{\theta_j}\}.$$

For a finite execution  $e = M \rightarrow_{p_1 \theta_1} M_1 \dots \rightarrow_{p_k \theta_k} M_k \in Exec_M^F$  starting from  $M$  and driven by a scheduler  $F$  we define

$$P_M^F(e) = p_1 \cdot \llbracket M'_1 \rrbracket_{\theta_1}(M_1) \cdot \dots \cdot p_k \cdot \llbracket M'_k \rrbracket_{\theta_k}(M_k)$$

where  $\forall j \leq k$ ,  $p_j = F(M_{j-1})(\llbracket M'_j \rrbracket_{\theta_j})$ . We denote by  $last(e)$  the final state of a *finite* execution  $e$ , by  $e^j$  the prefix  $M \rightarrow_{p_1 \theta_1} M_1 \dots \rightarrow_{p_j \theta_j} M_j$  of length  $j$  of the execution  $e = M \rightarrow_{p_1 \theta_1} M_1 \dots \rightarrow_{p_j \theta_j} M_j \rightarrow_{p_{j+1} \theta_{j+1}} M_{j+1} \dots$ , and by  $e \uparrow$  the set of  $\bar{e}$  such that  $e \leq_{prefix} \bar{e}$ . We write  $M \rightarrow_*^F M'$  if there exists a finite execution  $e \in Exec_M^F$  with  $last(e) = M'$ .

We define the probability space on the executions starting from a given network  $M$  as follows. Given a scheduler  $F$ ,  $\sigma Field_M^F$  is the smallest sigma field on  $Exec_M^F$  that contains the basic cylinders  $e \uparrow$ , where  $e \in Exec_M^F$ . The probability measure  $Prob_M^F$  is the unique measure on  $\sigma Field_M^F$  such that  $Prob_M^F(e \uparrow) = P_M^F(e)$ . Given a measurable set of networks  $H$ , we denote by  $Exec_M^F(H)$  the set of executions starting from  $M$  and crossing a state in  $H$ . Formally  $Exec_M^F(H) = \{e \in Exec_M^F \mid last(e^j) \in H \text{ for some } j\}$ . We denote the probability for a network  $M$  to evolve into a network in  $H$ , according to the policy given by  $F$ , as  $Prob_M^F(H) = Prob_M^F(Exec_M^F(H))$ . Note that the use of probabilistic schedulers allows us to model networks as discrete time Markov chains (DTMCs). This is the result of the application of a *two level* probability distribution: the reduction semantics maps a network  $M$  into a probability distribution in the set  $behave(M)$  while, in turn, the probabilistic scheduler maps  $M$  into a probability distribution  $\phi$  over the probability distributions in the set  $behave(M)$ , giving rise to a fully probabilistic model.

The notion of barb introduced below denotes an observable transmission with a certain probability according to a fixed scheduler. We first introduce a notion of *strong barb*: for a network  $M$ , we write  $M \downarrow_K$  when  $M \equiv [(\tilde{v})_{L,r}.P]_l^J | M'$  with  $\emptyset \neq K \subseteq L$  and for all  $k \in K$ ,  $d(l, k) \leq r$ . Roughly, a transmission is observable only if at least one location in the set of the intended recipients is able to receive the message. We say that a network  $M$  has a *barb* with probability  $p$  at the set  $K$  of locations, according to the scheduler  $F$ , written  $M \downarrow_p^F K$ , if  $Prob_M^F(\{M' \mid M \rightarrow_*^F M' \downarrow_K\}) = p$ . Intuitively, for a given network  $M$  and scheduler  $F$ , if  $M \downarrow_p^F K$  then there is a positive probability that  $M$ , driven by  $F$ , performs a transmission and at least one of the intended recipients is able to correctly listen to it. In the following, we introduce a notion of probabilistic observational congruence relative to a specific set of schedulers  $\mathcal{F} \in Sched$ . Since our semantics is contextual, we need to ensure that the set of schedulers we consider allows the specific networks we analyse to interact with any possible context. Hence for a set  $\mathcal{F}$  of schedulers we define the *contextual superset*  $\mathcal{F}_c$  of  $\mathcal{F}$ , as the largest set of schedulers allowing networks to interact with any pos-

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(Output) $\frac{-}{\langle \tilde{v} \rangle_{L,r}.P \xrightarrow{\tilde{v}_{L,r}} P}$	(Input) $\frac{-}{(\tilde{x}).P \xrightarrow{\tilde{v}} P\{\tilde{v}/\tilde{x}\}}$	
(Then) $\frac{P \xrightarrow{\eta} P'}{[\tilde{v} = \tilde{v}]P, Q \xrightarrow{\eta} P'}$	(Else) $\frac{Q \xrightarrow{\eta} Q' \quad \tilde{v}_1 \neq \tilde{v}_2}{[\tilde{v}_1 = \tilde{v}_2]P, Q \xrightarrow{\eta} Q'}$	(Rec) $\frac{P\{\tilde{v}/\tilde{x}\} \xrightarrow{\eta} P'}{A\langle \tilde{v} \rangle \xrightarrow{\eta} P'} \quad A\langle \tilde{v} \rangle \stackrel{\text{def}}{=} P$

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**Table 3: LTS rules for Processes**

sible context even when driven by  $\mathcal{F}$  (see [4] for a formal definition). It holds that  $Sched_{\mathcal{C}} = Sched$ . Hereafter, a context  $C[\cdot]$  is a term with a hole defined by the grammar:  $C[\cdot] ::= [\cdot] \mid [\cdot]M \mid M[\cdot]$ . Our probabilistic observational congruence relative to a specific set of schedulers is defined as follows.

**DEFINITION 2.** *Given a set  $\mathcal{F} \subseteq Sched$  and a relation  $\mathcal{R}$  over networks:*

- $\mathcal{R}$  is barb preserving relative to  $\mathcal{F}$  if  $MRN$  and  $M \Downarrow_p^F K$  for some  $F \in \mathcal{F}_{\mathcal{C}}$  implies that  $\exists F' \in \mathcal{F}_{\mathcal{C}}$  such that  $N \Downarrow_p^{F'} K$ .
- $\mathcal{R}$  is reduction closed relative to  $\mathcal{F}$  if  $MRN$  implies that  $\forall F \in \mathcal{F}_{\mathcal{C}} \exists F' \in \mathcal{F}_{\mathcal{C}}$  such that  $\forall C \in \mathcal{N}/\mathcal{R}$ ,  $Prob_M^F(C) = Prob_N^{F'}(C)$ .
- $\mathcal{R}$  is contextual if  $MRN$  implies that  $C[M] \mathcal{R} C[N]$  for every context  $C[\cdot]$ .
- Probabilistic observational congruence relative to  $\mathcal{F}$ , written  $\cong_p^{\mathcal{F}}$ , is the largest symmetric relation over networks which is reduction closed, barb preserving and contextual.

Two networks are related by  $\cong_p^{\mathcal{F}}$  if they exhibit the same probabilistic (connectivity) behaviour relative to  $\mathcal{F}$ . In the next section a bisimulation-based proof technique for  $\cong_p^{\mathcal{F}}$  is developed in order to provide an efficient method to check whether two networks are related by  $\cong_p^{\mathcal{F}}$ .

*Deciding the Observational Congruence.* We express the semantics of the calculus in terms of labelled transition systems (LTS) which are built upon two sets of rules: one for processes and one for networks. Table 3 presents the LTS rules for processes. Transitions are of the form  $P \xrightarrow{\eta} P'$ , where  $\eta$  ranges over input and output actions:  $\eta ::= \tilde{v} | \tilde{v}_{L,r}$ .

Table 4 presents the LTS rules for networks. Transitions are of the form  $M \xrightarrow{\gamma} \llbracket M' \rrbracket_{\theta}$ , where  $M$  is a network and  $\llbracket M' \rrbracket_{\theta}$  is a distribution over networks. Probabilities are used to model the mobility of nodes. Tag  $\gamma$  ranges over the labels:

$$\gamma ::= L!\tilde{v}[l,r] \mid ?\tilde{v}@l \mid R!\tilde{v}@K \mid \tau.$$

Rule (Snd) models the sending of tuple  $\tilde{v}$  to a specific set  $L$  of locations with transmission radius  $r$ , while rule (Rcv) models the reception of  $\tilde{v}$  at  $l$ . Rule (Bcast) models the broadcast message propagation: all the nodes lying within the transmission cell of the sender may receive the message, regardless of the fact that they lie in one of the locations in  $L$ . Rule (Obs) models the observability of a transmission: every transmission may be detected (and hence *observed*) by any recipient lying in one of the observation locations within the transmission cell of the sender. The label  $R!\tilde{v}@K$  represents the transmission of the tuple  $\tilde{v}$  of messages: the set  $R$  is the set of all the locations receiving the message, while its subset  $K$  contains only the locations where the transmission is observed. Rule (Lose) models message loss.

As usual,  $\tau$ -transitions denote non-observable actions. Rule (Move) models node mobility according to the probability distribution  $\mu_l^1$ . Finally, (Par) is standard.

Based on the LTS semantics, we define a probabilistic labelled bisimilarity that is a characterisation of our *probabilistic observational congruence*. It is built upon the actions:

$$\alpha ::= ?\tilde{v}@l \mid R!\tilde{v}@K \mid \tau.$$

We write  $lbehave(M)$  for the set of all possible behaviours of  $M$ , that is  $lbehave(M) = \{(\alpha, \llbracket M' \rrbracket_{\theta}) \mid M \xrightarrow{\alpha} \llbracket M' \rrbracket_{\theta}\}$ . A scheduler<sup>1</sup> for the labelled semantics is a function  $F$  assigning a probability to each pair  $(\alpha, \llbracket M \rrbracket_{\theta}) \in lbehave(M)$  with a network  $M$ . We denote by  $LSched$  the set of schedulers for the LTS semantics. A labelled *execution*  $e$  of a network  $M$  driven by a scheduler  $F$  is a finite (or infinite) sequence of steps:  $M \xrightarrow{\alpha_1}_{p_1\theta_1} M_1 \xrightarrow{\alpha_2}_{p_2\theta_2} M_2 \dots \xrightarrow{\alpha_k}_{p_k\theta_k} M_k$ . With abuse of notation, we define  $Exec_M^F$ ,  $last(e)$ ,  $e^j$  and  $e \uparrow$  as for unlabeled executions.

Since we are interested in weak observational equivalences, that abstract over  $\tau$ -actions, we introduce the notion of *weak action* as follows:  $\Longrightarrow$  is the transitive and reflexive closure of  $\xrightarrow{\tau}$ ;  $\xrightarrow{\alpha}$  denotes  $\Longrightarrow \xrightarrow{\alpha} \Longrightarrow \forall \alpha \neq \tau$ . We write  $\xrightarrow{\alpha}$  for the weak action  $\xrightarrow{\alpha}$  if  $\alpha \neq \tau$ , and  $\Longrightarrow$  otherwise. We denote by  $Exec_M^F(\xrightarrow{\alpha}, H)$  the set of all executions that, starting from  $M$ , according to the scheduler  $F$ , lead to a network in the set  $H$  by performing  $\xrightarrow{\alpha}$ . We define the probability of reaching a network in  $H$  from  $M$  by performing  $\xrightarrow{\alpha}$ , according to a scheduler  $F$  as  $Prob_M^F(\xrightarrow{\alpha}, H) = Prob_M^F(Exec_M^F(\xrightarrow{\alpha}, H))$ . For  $\mathcal{F} \subseteq Sched$ , we denote by  $\tilde{\mathcal{F}}_{\mathcal{C}} \subseteq LSched$  its *contextual superset* for the LTS semantics (see [4]).

**DEFINITION 3.** *Let  $M$  and  $N$  be two networks. An equivalence relation  $\mathcal{R}$  over networks is a probabilistic labelled bisimulation relative to a set  $\mathcal{F}$  of schedulers, if  $MRN$  implies: for all schedulers  $F \in \tilde{\mathcal{F}}_{\mathcal{C}}$  there exists a scheduler  $F' \in \tilde{\mathcal{F}}_{\mathcal{C}}$  such that for all  $\alpha$  and for all classes  $C \in \mathcal{N}/\mathcal{R}$ :*

- if  $\alpha \neq ?\tilde{v}@l$  then  $Prob_M^F(\xrightarrow{\alpha}, C) = Prob_N^{F'}(\xrightarrow{\alpha}, C)$ ;
- if  $\alpha = ?\tilde{v}@l$  then either  $Prob_M^F(\xrightarrow{\alpha}, C) = Prob_N^{F'}(\xrightarrow{\alpha}, C)$  or  $Prob_M^F(\xrightarrow{\alpha}, C) = Prob_N^{F'}(\Longrightarrow, C)$ .

Probabilistic labelled bisimilarity relative to  $\mathcal{F}$ , written  $\approx_p^{\mathcal{F}}$ , is the largest probabilistic labelled bisimulation relative to  $\mathcal{F}$  over networks.

Probabilistic labelled bisimilarity is a characterization of our probabilistic observational congruence [4].

**THEOREM 1.**  $M \cong_p^{\mathcal{F}} N$  if and only if  $M \approx_p^{\mathcal{F}} N$ .

<sup>1</sup>We abuse notation and still use  $F$  to denote a scheduler for the LTS semantics.

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$\text{(Snd)} \frac{P \xrightarrow{\tilde{v}_{L,r}} P'}{[P]_l^J \xrightarrow{L\tilde{v}[l,r]} \llbracket [P']_l^J \rrbracket_\Delta}$	$\text{(Rcv)} \frac{P \xrightarrow{\tilde{v}} P'}{[P]_l^J \xrightarrow{?\tilde{v}@l} \llbracket [P']_l^J \rrbracket_\Delta}$	
$\text{(Bcast)} \frac{M \xrightarrow{L\tilde{v}[l,r]} \llbracket M' \rrbracket_\Delta \quad N \xrightarrow{?\tilde{v}@l'} \llbracket N' \rrbracket_\Delta}{M N \xrightarrow{L\tilde{v}[l,r]} \llbracket M' N' \rrbracket_\Delta} \quad d(l, l') \leq r$		
$\text{(Obs)} \frac{M \xrightarrow{L\tilde{v}[l,r]} \llbracket M' \rrbracket_\Delta}{M \xrightarrow{R\tilde{v}@K} \llbracket M' \rrbracket_\Delta} \quad R \subseteq \{l' \in \text{Loc} : d(l, l') \leq r\} \quad K = R \cap L, \quad K \neq \emptyset$		
$\text{(Lose)} \frac{M \xrightarrow{L\tilde{v}[l,r]} \llbracket M' \rrbracket_\Delta}{M \xrightarrow{\tau} \llbracket M' \rrbracket_\Delta}$	$\text{(Move)} \frac{}{[P]_l^J \xrightarrow{\tau} \llbracket [P]_l^J \rrbracket_{\mu_l^J}}$	$\text{(Par)} \frac{M \xrightarrow{\gamma} \llbracket M' \rrbracket_\theta}{M N \xrightarrow{\gamma} \llbracket M' N \rrbracket_\theta}$

---

Table 4: LTS rules for Networks

### 3. MEASURING ENERGY CONSUMPTION

In this section, based on the LTS semantics, we define a preorder over networks which allows us to study the performances, in terms of energy consumption, of different networks, but exhibiting the same connectivity behaviour. For this purpose we associate an energy cost with labelled transitions as follows. For a transmission with radius  $r$ , let

$$\text{En}(r) = \text{En}_{elec} \times \text{packet\_len} + \text{En}_{ampl} \times \text{packet\_len} \times r^2$$

where  $\text{En}_{elec}$  ( $nJ/b$ ) is the energy dissipated to run the transmitter circuit, while  $\text{En}_{ampl}$  ( $pJ/b/m^2$ ) is the radio amplifier energy (see [11]). We define

$$\text{Cost}(M, N) = \begin{cases} \text{En}(r) & \text{if } M \xrightarrow{L\tilde{v}[l,r]}_\Delta N \\ & \text{for some } L, \tilde{v}, l \text{ and } r \\ 0 & \text{otherwise} \end{cases}$$

For an execution  $e = M_0 \xrightarrow{\alpha_1}_{\theta_1} M_1 \xrightarrow{\alpha_2}_{\theta_2} M_2 \dots \xrightarrow{\alpha_k}_{\theta_k} M_k$ ,

$$\text{Cost}(e) = \sum_{i=1}^k \text{Cost}(M_{i-1}, M_i).$$

Let  $H$  be a set of networks; we denote by  $\text{Paths}_M^F(H)$  the set of all executions from  $M$  ending in  $H$  and driven by  $F$  which are not prefix of any other execution ending in  $H$ . More formally,  $\text{Paths}_M^F(H) = \{e \in \text{Exec}_M^F(H) \mid \text{last}(e) \in H \text{ and } \forall e' \text{ such that } e <_{\text{prefix}} e', e' \notin \text{Paths}_M^F(H)\}$ .

DEFINITION 4. *The average cost of reaching a set of networks  $H$  from an initial network  $M$  according to the scheduler  $F$  is*

$$\text{Cost}_M^F(H) = \frac{\sum_{e \in \text{Paths}_M^F(H)} \text{Cost}(e) \times P_M^F(e)}{\sum_{e \in \text{Paths}_M^F(H)} P_M^F(e)}.$$

The average cost is computed by weighting the cost of each execution by its probability according to  $F$  and normalized by the overall probability of reaching  $H$ .

DEFINITION 5. *Let  $\mathcal{H}$  be a countable set of sets of networks and let  $\mathcal{F} \subseteq \text{LSched}$  a set of schedulers. We write*

$$N \sqsubseteq_{\mathcal{H}}^{\mathcal{F}} M,$$

*if  $N \approx_p^{\mathcal{F}} M$  and,  $\forall$  schedulers  $F \in \text{LSched}$  and  $\forall H \in \mathcal{H}$ ,  $\exists$  a scheduler  $F' \in \text{LSched}$  such that  $\text{Cost}_N^{F'}(H) \leq \text{Cost}_M^F(H)$ .*

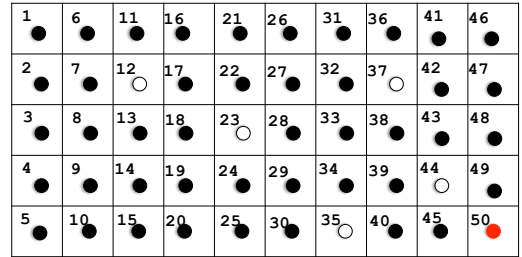


Figure 1: Topology of the Static Network (SN)

### 4. STUDYING GOSSIP PROTOCOLS

Gossip protocols are a family of communication protocols inspired by the way that gossiping disseminates information in social networks. A gossip protocol is a variant of the flooding algorithm, where each node forwards a message with some probability to reduce the overhead of the routing protocols. Gossiping based routing protocols are commonly used in large-scale networks (see, e.g., [3, 10, 7]) to reduce the number of retransmissions and the energy cost. In this section we show that our framework is suitable for providing an integrated automatic analysis of the gossip strategy in terms of both connectivity maintenance and energy consumption. In particular, we assume that, when a node receives a message, it forwards it with a fixed probability  $\text{psend}$  and discards it with probability  $1 - \text{psend}$ . Common values for  $\text{psend}$  ranges from 0.6 to 0.8: it is shown that, in practical scenarios, these values provide a reduction of more than 30% of the forwarding transmissions without deteriorating the network connectivity [7]. Here we consider two different network configurations on a rectangular area of  $50 \times 100\text{m}$ . We assume omnidirectional antenna and a fixed transmission power for each sensor node, which covers circular areas with a radius of 10m. In the following, we denote by  $[P_i]_l^J$  the sensor node  $i$  located at  $l$ , executing the process  $P_i$  and moving according to the transition matrix  $\mathbf{J}$ . We study the behaviour of the networks by varying the value of the parameter  $\text{psend}$ . The first network we consider consists of 50 static nodes, evenly distributed within the network area (see Figure 1). Node mobility is characterised by the identity matrix  $\mathbf{I}$ . In our tests, we consider a

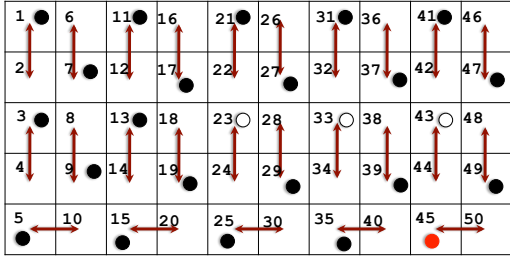


Figure 2: Topology of the Mobile Network (MN)

fixed receiver  $[P_{50}]_{50}^I$ , while the sender node's location varies in the set  $\{12, 23, 35, 37, 44\}$ , in order to study how the connectivity behaviour of the network changes, depending on the distance between the sender node and the receiver. The network is expressed by the term:

$$M_j \stackrel{\text{def}}{=} \sum_{i=1}^{50} [P_i]_i^I,$$

with  $j \in \{12, 23, 35, 37, 44\}$ , and  $P_i \stackrel{\text{def}}{=} (x_i) \cdot \langle x_i \rangle_{\{50\}, 10} \cdot P_i$ ,  $\forall i \notin \{j, 50\}$ ,  $P_j \stackrel{\text{def}}{=} \langle x_j \rangle_{\{50\}, 10} \cdot P_j$ ,  $P_{50} \stackrel{\text{def}}{=} (x_{50}) \cdot P_{50}$ , modelling the communication between  $[P_j]_j^I$  and  $[P_{50}]_{50}^I$ . The second configuration consists of 25 mobile sensor nodes again evenly distributed within the network area. Each sensor node can move between two adjacent locations, modelling the instability caused by, e.g., environmental conditions (see Figure 2). The probability distribution associated with node mobility can be captured by the transition matrix  $\mathbf{J}$  such that:  $\mathbf{J}_{l(l+5)} = \mathbf{J}_{(l+5)l} = \varepsilon \forall l \in \{5, 15, 25, 35, 45\}$ , and  $\mathbf{J}_{l(l+1)} = \mathbf{J}_{(l+1)l} = \varepsilon$  for all the other odd locations in the network area, and  $\mathbf{J}_{ll} = 1 - \varepsilon$  for all the locations, with  $0 < \varepsilon < 1$ . Notice that the choice of  $\varepsilon$  and the definition of the scheduler allow us to model the relative speed between movements and transmissions. Henceforth we assume that  $\varepsilon = 0.8$ . This network is expressed by the term:

$$N_h \equiv \sum_{i=1}^{25} [P_i]_{(2i-1)}^J,$$

with  $h \in \{12, 17, 22\}$ , where  $P_i \equiv (x_i) \cdot \langle x_i \rangle_{\{45, 50\}, 10} \cdot P_i$ ,  $\forall i \notin \{h, 25\}$ ,  $P_h \equiv \langle x_h \rangle_{\{45, 50\}, 10} \cdot P_h$  and  $P_{25} \equiv (x_{25}) \cdot P_{25}$ , modelling the communication between  $[P_h]_{(2h-1)}^J$  and  $[P_{25}]_{25}^J$ , with  $z \in \{45, 50\}$  the set of locations where we expect to find  $P_{25}$ . We model several different gossip strategies by varying the value of  $\mathbf{psend}$  in the interval  $[0.6 - 1.0]$ . In particular, for each value of  $\mathbf{psend}$  we assume a set  $\mathcal{F}_{\mathbf{psend}}$  of schedulers such that, at each step, the probability for each node to perform a synchronisation or a movement is the same. Moreover, we do not consider message loss due to link failure or other environmental causes: a message can be lost only when a node discards it, consistently with the protocol. The analysis is performed using the PRISM model checker [8] (see the Appendix for details). The first step of our methodology consists in translating the process algebraic definition of our networks into the language supported by PRISM. This can be achieved in a purely algorithmic way. In general the exact analyses of real WSNs' models is unfeasible due to the explosion of the state space of the model. For this reason, we choose to perform an exact analysis to study problems of equivalences or performances in case of small components, e.g., to replace a network's node with a functionally equivalent one that has better performances in terms of through-

put or energy consumption. Conversely, when studying the overall properties of wide WSNs we apply the *approximate model checking* (also known as *statistical model checking*), that relies on a Monte Carlo simulation of the underlying DTMC. As a consequence, PRISM will compute estimates of the desired indices rather than results, whose precision is controlled by means of confidence interval specifications (absolute width and confidence). When simulation is adopted, the estimates are obtained by sampling, i.e., generating a large number of random paths through the process underlying the model, hence avoiding the generation of whole DTMC. In our case studies we assume that the sender node keeps retransmitting the same packet until the destination node receives it. The outcomes of this study allow us to determine the expected number of retransmissions of the same packet that are needed to reach the intended recipient or, in more detail, which is the number of retransmissions needed in order to reach the destination with a probability higher than a given threshold. Our goal is the comparison between the different network configurations, according to the definition of energy aware preorder introduced in Section 3. Concerning the termination of the simulations, the next proposition states that, by varying the sender location, the packet eventually reaches the intended recipient.

PROPOSITION 1.

- (i)  $\forall \mathcal{F}_{\mathbf{psend}}, \mathbf{psend} \in [0.6 - 1.0]$  and  $\forall j_1, j_2 \in \{12, 23, 35, 37, 44\}$

$$M_{j_1} \approx_p^{\mathcal{F}_{\mathbf{psend}}} M_{j_2}.$$

- (ii)  $\forall \mathcal{F}_{\mathbf{psend}}, \mathbf{psend} \in [0.6 - 1.0]$  and  $\forall h_1, h_2 \in \{12, 17, 22\}$

$$N_{h_1} \approx_p^{\mathcal{F}_{\mathbf{psend}}} N_{h_2}.$$

Once we proved that the networks we are considering have the same connectivity, we are ready to compare their energy costs, by changing the value of  $\mathbf{psend}$  and the distance among the sender and the receiver. Using the PRISM model checker, we exploit the possibility of defining reward measures to compute the energy cost function defined in Section 3. Assuming that the energy spent for each transmission is fixed and that all the nodes have the same physical characteristics, we simply count the number of transmissions rather than summing their energy cost. The cost function is expressed in terms of a Probabilistic Computation Tree Logic (PCTL) formula in the PRISM property specification language, augmented with *rewards* (or costs), which are real-valued quantities associated with states and/or transitions (see the Appendix for more details). Specifically, we verify the formula  $\mathbf{R} = ? [\mathbf{F} \text{ goal}]$ , which expresses the cumulative expected energy cost to complete the communication.

*Validation of the simulator.* We have validated our simulator by comparing our estimates with those obtained in [7] (see Figure 3 and 4).

*Simulation of static Networks.* The estimates for the static network are shown in Figure 5. The simulations have been performed with an average of 10000 experiments, a maximum confidence interval width of 1% of the estimated measure based on 95% of confidence. The plots show how the distance between sender and receiver critically influences the energy performance of the algorithm. For a distance larger than 30m we have a monotonic decreasing plot showing that, for large distances, the gossip protocol can cause energy waste. Using the standard flooding strategy ( $\mathbf{psend} = 1.0$ )

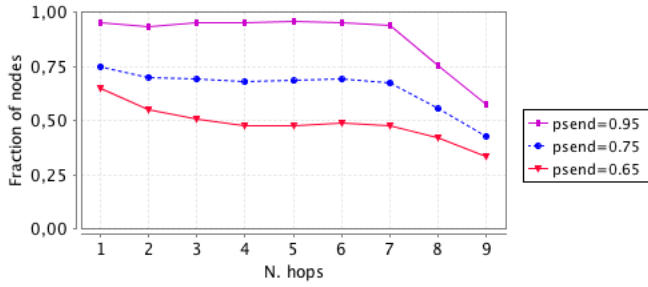


Figure 3: (SN) Fraction of nodes reached by a transmission

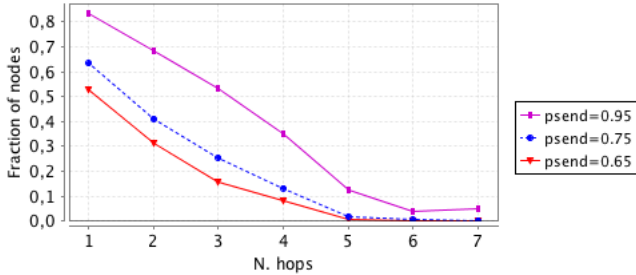


Figure 4: (MN) Fraction of nodes reached by a transmission

all the cases converge to 49, because each node will forward the message exactly one time. We can verify that there exists a preorder among different network configurations within the confidence of the simulation.

**PROPOSITION 2.** For all  $\mathbf{psend} \in \{0.6, 0.65, 0.7, 0.75, 0.8, 0.85, 0.9, 1.0\}$  and for all  $j_1 < j_2 \in \{12, 23, 35, 37, 44\}$  it holds that  $M_{j_1} \sqsubseteq_{\mathcal{H}}^{\mathcal{F}_{\mathbf{psend}}} M_{j_2}$ , where  $\mathcal{H}$  is the set of network configurations where the communication has been successfully completed.

Figure 6 shows the expected number of retransmissions that the sender node must perform before the communication is successfully completed. Notice that the smaller is the value of  $\mathbf{psend}$ , the higher is the probability that the message is lost during the path, forcing a new transmission (for the sake of simplicity we don't model the acknowledges, but we assume that the sender node will wait for an acknowledge until a timeout occurs, then it will transmit again); hence, even if a small value of  $\mathbf{psend}$  reduces the forwarding explosion, it may increase the number of replications.

*Simulation of networks with mobility.* Figure 7 shows the estimates of the expected energy cost for a successful transmission in the WSN with mobility.

The mobility of the nodes critically increases the size of the state space, hence the obtained results have wider confidence intervals than those of the static network simulation, based on 95% of confidence. However, the results are very similar to the previous case: for distances larger than 25m the gossip protocol causes a very high energy waste.

**PROPOSITION 3.** For all  $\mathbf{psend} \in \{0.6, 0.65, 0.7, 0.75, 0.8, 0.85, 0.9, 1.0\}$  and for all  $h_1 < h_2 \in \{12, 17, 22\}$ , it holds that  $N_{h_1} \sqsubseteq_{\mathcal{H}}^{\mathcal{F}_{\mathbf{psend}}} N_{h_2}$  where  $\mathcal{H}$  is the set of network configurations where the communication has been successfully completed.

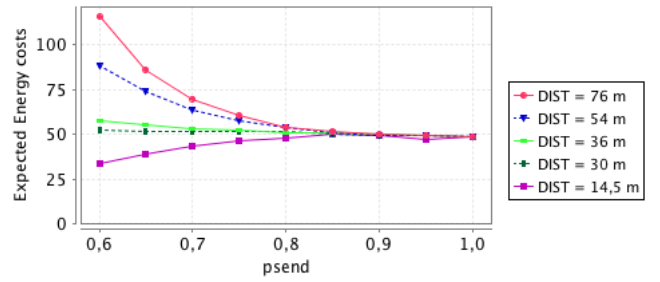


Figure 5: (SN) Expected energy cost

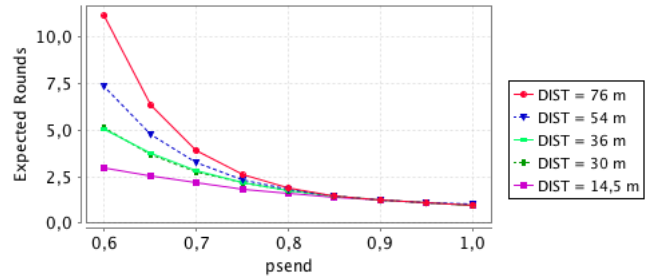


Figure 6: (SN) Expected number of transmissions for a successful communication

figures where the communications has been successfully completed.

Figure 8 shows the average number of retransmissions that the sender must perform before the communication has successfully completed.

## 5. RELATED WORK AND CONCLUSIONS

A large amount of research on sensor networks has been recently reported in the last decade. Several papers address the problem of studying the energy consumption for specific communication protocols. For instance, in [18] the authors define a Markov reward process (see, e.g., [14]) modelling some protocols for point to point reliable transmissions. A steady-state quantitative analysis is then obtained, and from this the average performance indices are computed. In [1], Bernardo et al. present a methodology to predict the impact of power management techniques on system functionality and performance. In [16], the authors define a set of metrics to analyse the energy consumption which are then estimated through simulation, and show how some modifications in the protocols can improve their efficiency. In [7], gossip protocols running on WSNs are studied but the authors develop an ad hoc simulator to estimate their performances. Conversely, in our setting, a general purpose tool, e.g., PRISM, can be used since the performance indices or properties to be evaluated (or estimated) can be formally specified according to a rigorous logic. Moreover, with respect to all the above mentioned contributions, the model we propose here aims at providing a common framework for automatically perform both qualitative and quantitative analyses. The energy preorder defined in Section 3 can be efficiently decided for small networks' components using model checking methods and hence one may decide to replace a node with

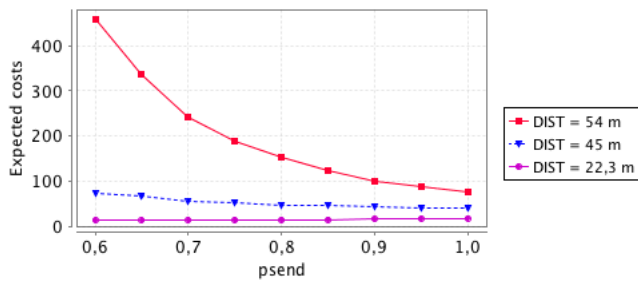


Figure 7: (MN) Expected energy cost

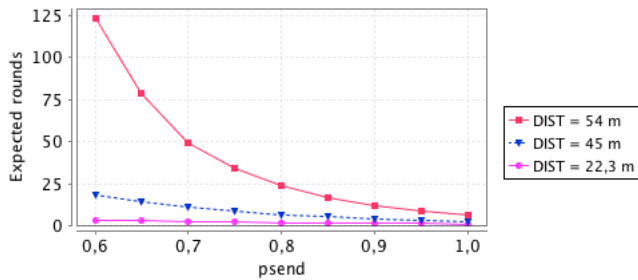


Figure 8: (MN) Expected number of transmissions for a successful communication

another which is behaviourally equivalent but less energy consuming; conversely, when the complexity of the process underlying the model makes exact analyses unfeasible, approximate (or statistical) model checking can be employed. This corresponds to the well-known Monte Carlo simulation; using the temporal logic implemented in the tool, one can verify a proposition within a certain level of confidence (e.g., the network equipped with a certain protocol is connected with a confidence of 99.9%). To the best of our knowledge such a qualitative and quantitative approach supported by the same tool represents a novelty in the study of WSNs.

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